

A Linear Programming Approach to Determine an Optimum Value in a Single Currency Project of West Africa**

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ABSTRACT

The paper discusses the primary and secondary convergence conditions for the second monetary zone in West Africa. The focal point is however the primary conditions as these provide the basis for the attainment of the secondary conditions. Panel data for the research are obtained from the West African Monetary Agency website: www.wami.imao.org. The variables are those given in the primary conditions and these are first tested for unit root and stationarity for each country. A panel cointegration test is then applied to obtain a long-run equation which is used as an objective function in the Simplex method of linear programming with the primary conditions as constraints. The panel unit root test results show that all the variables are integrated with the degree of integration varying from zero to one for different countries. The stationarity test confirms the result, as the variables are non-stationarity in level for some countries but stationarity for others. Since the unit root and stationarity test show conflicting results, the pooled mean group estimator is used to obtain long-run cointegration equation. This equation can be applied irrespective of whether the variables are integrated or not. Linear programming is then used to obtain the optimal condition for attainment of a single currency project for West Africa. The result shows that the objective function is minima at 0.0462 with inflation contributing more to the variation in the government external reserves. The paper recommend that Central Banks in those countries preparing for second monetary zone should avoid implementing inflation targeting as a way to solving their economic problem.

ARTICLE INFO

Keywords: Simplex method; optimal condition; convergence criteria; objective function; pooled group mean estimator; inflation targeting.

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Article Submitted 13-01-2014
Article Accepted 28-02-2014

**Article previously published in
EJEM 2014, vol 1, No. 1

1. INTRODUCTION

Regional economic groupings of countries to form monetary unions are fast becoming the norm in the world today. Many regions including those in Africa (in particular West Africa) have set-up one time table or the other towards a single monetary union. The Economic Community of West African States (ECOWAS) is a regional body which was formed on May 28, 1975 with fourteen members. Countries that signed the treaty creating ECOWAS include the following: Benin

Republic, Burkina Faso, Gambia, Ghana, Guinea Bissau, Guinea Conakry, Ivory Coast, Liberia, Mali, Niger, Nigeria, Senegal, Sierra Leone and Togo. Out of these countries, eight of them, Benin Republic, Burkina Faso, Guinea Bissau, Ivory Coast, Mali, Niger, Senegal and Togo already have a single monetary union, Union Économique et Monétaire Ouest-Africaine (UEMOA) using a single currency, franc de la Communauté Financière de l'Afrique (CFA franc), and a common central bank, the Banque Centrale des Etats de l'Afrique de Ouest (BCEAO) (Benassy-Quere and Coudet, 2005).

The issue of a single monetary zone for the West African sub-region was on the initial agenda of the ECOWAS members at its inception in 1975 (Masson and Pattillo, 2001). They formulated the following primary objectives as targets to be achieved:

- (i) Integration of monetary and fiscal policies to eventually create a common market,
- (ii) Creation of a single currency,
- (iii) Improve welfare of members through intra-regional trade (see Masson and Pattillo, 2001 & Ogunkola, 2005 for details).

The initial enthusiasm of members made them to pursue these objectives with vigor (Ojo, 2003). The lack of economic power however made the volume of intra-regional trade to be very low (Itsede, 2002 and Ogunkola, 2005). Similarly the objective to form a single monetary union for the entire region is still far away from being realized (Benassy-Quere and Coudet, 2005). Instead the five Anglophone countries Ghana, Nigeria, The Gambia, Guinea and Sierra Leone have gone ahead to initiate another monetary union tagged "The West African Monetary Zone (WAMZ)" which is expected to merge with the UEMOA after its take-off on December 1, 2015. Till date there is no assurance that this arrangement for the five countries will work.

Four primary and six secondary convergence criteria have been agreed upon. The targets and schedule of primary and secondary convergence criteria are shown in Table 1 and Table 2 below.

Table 1: Targets for Primary Convergence Criteria

Criteria	Benchmark
Inflation Rate	Single digit \leq (10%)
Gross External reserves (Months of Imports)	\geq 3 months
Central Bank Financing of Fiscal deficit as % of Previous Tax's Revenue	\leq 10%
Fiscal Deficit/Surplus/GDP (%) Excl. Grants	\leq 4%

Table 2: Targets for Secondary Convergence Criteria

Criteria	Benchmark
Change in Domestic Arrears	$\leq 0\%$
Tax Revenue/GDP ratio	$\geq 20\%$
Salary Mass /Total Tax Revenue	$\leq 35\%$
Exchange Rate (+:Depreciation/ -: Appreciation)	$\leq \pm 15\%$
Real Interest Rate	$> 0\%$
Domestically Financed Investments/Domestic Revenue	$> 20\%$

The core issues under the ECOWAS monetary union are the convergence in microeconomic variables and stability of exchange rate as stated in the primary and secondary convergence criteria. In addition to facilitating policy coordination, it affords the opportunity to ensure microeconomic stability (WAMA, 2009). Available literature (WAMA, 2009 & Egwaikhide and Ogunleye, 2010) shows that as at end of 2008 three primary criteria were satisfied except inflation whose average was 13.6 above the targeted value ($\leq 10\%$). It is sad to note that from that point, these countries could not maintain the momentum for which they used to satisfy these primary conditions. As demonstrated above, all the countries performed poorly in attaining the targets set by the monetary agency. The five countries have to tighten their belts if not, the dream of a single currency come December 2015 will be a mirage.

The past had not been pleasant for ECOWAS, in terms of implementation of set goals. The objectives set by the regional body have not been achieved on time. The region remains amongst the poorest in the world, with weak institutions as reported in Benassy-Quere and Coudet (2005) and Roudet et al. (2007). The big question is, "can the goals set by ECOWAS be achieved in the near future?"

Many solutions have been suggested in the literature as documented by Ojo (2003), Itsede (2002), Obadan (2002), Ogunkola (2002), Masson and Pattillo (2001, 2003), Benassy-Quere and Coudet (2005), Ogunkola (2002, 2005) and WAMI (2006, 2007). Masson and Pattillo (2001) for instance argued in favour of genuine trade liberalization as a precondition for successful monetary unification with an independent single currency. Quoting from Hanink and Owusun (1998), and Yeats (1998), they observed that because of low market potentials, high transportation costs, similar factor endowments, liberalization might not be the solution. Itsede (2002) quoting from Corden (1972) emphasized the concept of monetary union as premised on two essential but interwoven parts. These include: exchange rate union which is an area in which exchange rates bear permanent relationship to each other and currency convertibility where exchange rate controls are permanently removed. Monetary union is then defined as the existence of a single

monetary zone with a high degree of monetary stability in furtherance of economic integration (Itsede, 2002).

Other obstacles besides the conditions laid in the OCA¹ theory that militate against monetary integration all over the world are lack of commitment, political will and strong leadership to implement the policies and actions that will move the programmes forward (Itsede, 2002). In West Africa, another factor is the non convertibility of the currencies of countries in the zone. Besides the UEMOA countries that already use the CFA Franc whose convertibility is backed by the French Treasury, all the other currencies (including the Nigerian naira) are not yet fully convertible. According to Obadan (2002), for a monetary union to be successful, the several currencies in the integrating zone must be fully convertible one into the other at permanently fixed exchange rates, thereby effectively creating a single currency. As stated earlier, the focus of the paper is to use the convergence criteria given at Table 1 to obtain an optimum convergence point for these economic variables. This point is to serve as a point of convergence to be satisfied by the primary conditions, hence this study.

2. THEORETICAL FRAMEWORK

Before the commencement of monetary union, conditions for sustenance have to be in place. Implementing the monetary union before these primary and secondary convergence conditions are satisfied is like putting the cart before the horse, a major asymmetric shock would result in unbearable pressure within the union because of limited mobility and inadequate fiscal redistribution (Yuen, 2008). There is a common understanding that lasting economic convergence is a main panacea for a full monetary union. As such, significant degrees of convergence will have to be attained prior to monetary unification. The steps to be taken will be interdependent and will reinforce one another, in particular the development of monetary unification, must be based on sufficient progress in the field of convergence and then the unification of economic policies (Yuen, 2008).

The quantity theory of money explains the relationship between money supply and inflation. It states that there is direct relationship between quantity of money in an economy and prices of goods and services. In its simplest form the theory could be expressed as $MV=PQ$, where M is the money supply in an economy, V is the velocity of money in expenditures, P is the price level associated with

¹ An Optimum Currency Area (OCA) is an area that maintains fixed exchange rate within their members, and flexible exchange rate with trading partners outside the area (Appleyard and Field, 1998).

transactions for the economy and Q is the real output. The equation assumes that V and Q are constant in short terms. This has however been criticized by prominent economists like Keynes, especially the assumption that V is constant (Catao and Terrones, 2003).

The criticisms above suggest the relaxation of the assumptions underlying the quantity theory of money. The relaxation allows for a broader analysis of the possible interactions among the variables in the identity. Thus, given that the percentage change in a product, say U and V which is equal to the sum of the percentage changes ($\Delta U + \Delta V$). The quantity theory could be rewritten in terms of percentage changes as: $\Delta\pi + \Delta Q = \Delta M + \Delta V$. Rearrange the variables to get the identity:

$$\Delta M_t = \Delta\pi_t + \Delta Q_t - \Delta V_t \quad (1)$$

Equation (1) implies that an increase in money supply leads to increase in inflation or output. In other words, a decrease in money supply leads to decrease in inflation or output. Thus a decrease in velocity or a combination of changes in these variables (depending on a number of factors) such as elasticity of supply and the level of unemployment in the economy concerned leads to increase in money growth or decrease in inflation.

The relationship between fiscal deficit to GDP ratio and inflation has been investigated by many authors including Catao and Terrones (2003), Bhattacharya (2009) and Ammama et al. (2011). A study conducted by Catao and Terrones (2003) for 23 countries found that a 1% reduction in ratio of fiscal deficit to gross domestic product (GDP) ratio lower long-run inflation by 1.5 to 6%. They concluded that fiscal deficit is efficient determinant of inflation. Similarly, Bhattacharya (2009) tested the effect of fiscal deficit on inflation and concluded that any increase in demand caused by large deficit will raise prices. Simulation studies by Solomon and De-Wet (2004) indicate that inflation is very responsive to shocks in budget deficit as well as GDP. These studies did not investigate convergence, and the use the linear programming approach as adopted in this paper. Moreover, their focus was on long-run economic relationship between these variables.

3. MATERIALS AND METHODS

3.1. Data

The data for this research was collected from the West African Monetary Agency (WAMA) in the website <http://www.wami.imao.org>. The annual data is for 10 years 2001 - 2010. The variables are: inflation rate (π_t), ratio of fiscal deficit/surplus to GDP (b_t), Central Bank financing of fiscal deficit (d_t), and gross external reserves (m_t). These data were collected for the five countries The Gambia, Ghana, Guinea, Nigeria and Sierra Leone.

3.2. The Model Formulation

The starting point of the model formulation is the empirical relationship between fiscal deficit, gross external reserve (G) and inflation of a single equation model as given by Akçay et al. (2002), Vit (2003) and Solomon and De Wet (2004). Let us assume that inflation (π_t) ≤ 0.1 , gross external reserve (G) $\geq 0.25\text{Imp}_t$, where Imp_t is annual imports, central bank financing of fiscal deficit $\leq 0.1Tr_{t-1}$, where Tr_{t-1} is the tax revenue in the previous year and fiscal deficit to GDP ratio (Fd) ≤ 0.04 . The equation linking fiscal deficit to GDP ratio (Fd_t) with real interest rate (r_t) is given by Vit (2003) as:

$$r_t = c + Fd_t + \varepsilon_t \quad (2)$$

where c is a constant and ε_t is the random error.

However, similar to equation (1) Akçay et al. (2002) started with the following equation:

$$G_t - Tr_t + i_t B_{t-1} = \Delta M_t + \Delta B_t \quad (3)$$

where M_t is the reserve money, B_t is total stock of domestic and foreign debt, i_t is nominal interest rate, G_t and Tr_t are already explained above. If the budget deficit is inclusive of interest payment then

$$D_t^* = \Delta M_t + \Delta B_t \quad (4)$$

where D_t^* is the budget deficit. For countries with high inflation like those in WAMZ

$$d_t + b_{t-1}\rho_t = \Delta b_t \quad (5)$$

where d_t is the primary deficit less reserve money with each variable scaled by nominal output Y , b_t is total stock of domestic and foreign debt at period t , and ρ_t is reserve money (see Akcay et al. 2002 for details). Equation (4) therefore follows that:

$$b_{t-1} = \frac{1}{(1 + \rho_t)} (b_t - d_t) \quad (6)$$

The discounted debt-output ratio can be thus calculated using

$$X_t = b_t \prod_{k=1}^t (1 + \rho_k)^{-1} \quad (7)$$

Taking logarithm transformation of equation (6) and rearranging leads to

$$\log(b_t) = \log(X_t) + \sum_{k=1}^t (1 + \rho_k) \quad (8)$$

The starting point of the long run government budget constraint of Solomon and De Wet (2004) is:

$$\frac{B_{t-1}}{P_t} = \sum \frac{1}{r_j} \left(\tau_{t+j} - g_{t+j} + \left(M_{t+j} - \frac{M_{t-j-1}}{P_{t+j}} \right) \right) \quad (9)$$

where $\frac{B_{t-1}}{P_t}$ is government debt, r_j is discount rate, τ_t is the tax revenue, g_t is the government expenditure and M_t is the broad money supply. Catao and Terrones (2003) derived a long-run equation that explains the inflation rate by the budget deficit and money supply as:

$$\pi_t = \alpha b_t \frac{P_t}{M_t} \quad (10)$$

where α is the inverse linear multiplier, b_t is the budget deficit which is $b_t = g_t - \tau_t - B_{t-1}$ and $\frac{M_t}{P_t}$ is the money supply.

Studies carried out by Vit (2003) found the relationship between real interest rate (discount rate) and budget deficit as follows:

$$r_t = c + b_t + \varepsilon_t \quad (11)$$

where r_t is the real interest rate, b_t is the budget deficit, c and ε_t are drift and residual terms respectively. Considering equations (5), (9) and (10), we have the functional form of the model given as:

$$m_t = f(\pi_t, d_t, b_t) \quad (12)$$

Linearly, equation (11) can be expressed as

$$m_t = \phi_1 d_t + \phi_2 b_t + \phi_3 \pi_t + \varepsilon_t \quad (13)$$

where π_t is the inflation rate ≤ 0.1 , b_t is the fiscal deficit/surplus/GDP ratio ≤ 0.04 , d_t is Central Bank financing of fiscal deficit ≤ 0.1 , and m_t is the gross external reserves ≥ 0.25 .

The long-run economic relationships between these variables are obtained by conducting the cointegration test. The long run cointegration equation is:

$$\Delta y_t = \Pi y_{t-1} + \sum_{j=1}^{p-1} \Gamma_j \Delta y_{t-j} + x_t + \varepsilon_t \quad (14)$$

where $y_t = (m_t, d_t, b_t, \pi_t)'$ is a vector of endogenous variables, x_t is a vector of exogenous (explanatory) variables which include deterministic terms and ε_t is the residual error term.

This can be extended to the panel cointegration like the pooled mean group estimator of Persaran et al. (1999) given in section 3.4.

3.3. Panel Unit Root Tests

3.3.1 Im, Pesaran and Shin test

The Im, Pesaran and Shin (2003), test is based on N augmented Dickey-Fuller regressions:

$$\Delta y_{it} = \rho_i y_{i,t-1} + \sum_{j=1}^{p_i} \phi_{ij} \Delta y_{i,t-j} + \alpha_i + \gamma_i t + u_{it} \text{ for } i = 1, \dots, N; \quad t = 1, \dots, T \quad (15)$$

where T is the length of the sample, N is the cross-section dimension, $\Delta y_{i,t-j}$ are the lagged dependent variables included to eliminate serial correlation in the error term, p_i are the country-specific lag lengths, α_i and γ_i are country-specific intercepts (fixed effects) and trend parameters, respectively. The error term u_{it} is distributed as a white-noise random variable across i and t , with

$E(u_{it}) = 0$, $E(u_{it}^2) = \sigma_i^2 < \infty$ and $E(u_{it}^4) < \infty$. p_i , is assumed to be known and in practice it is estimated using an information criteria. The null hypothesis in the IPS test is $H_0 : \rho_i = 0$ for all i , which implies that all individual series contain a unit root (are independent random walks). The alternative hypotheses assume that $H_{1a} : \rho_1 < 0, \dots, \rho_{N_0} < 0$, and $H_{1b} : \rho_i = 0$ for $i = N_0 + 1, \dots, N$ (heterogeneous alternative). The IPS test averages the Augmented Dickey-Fuller statistics across groups to remove contemporaneous correlation in the errors. The statistic is given by:

$$\bar{t}_{NT} = \frac{1}{N} \sum_{i=1}^N t_{iT}(p_i, \varphi_i). \quad (16)$$

This statistic converges sequentially to normal distribution as $T \rightarrow \infty$ followed by $N \rightarrow \infty$, while the ratio $\frac{N}{T} \rightarrow \alpha > 0$, where α is a constant (see Lukacs, 1975). IPS compute the values $E[t_{iT}(p_i, \varphi_i)]$ and $Var[t_{iT}(p_i, \varphi_i)]$ and propose a statistic which tends asymptotically to the standard normal distribution.

$$W_T = \frac{\sqrt{N} \left[\bar{t}_{NT} - N^{-1} \sum_{i=1}^N E[t_{iT}(p_i, 0) | p_i = 0] \right]}{\sqrt{N^{-1} \sum_{i=1}^N Var[t_{iT}(p_i, 0) | p_i = 0]}} \xrightarrow[T, N \rightarrow \infty]{d} N(0, 1) \quad (17)$$

In practice, IPS test has been recommended for use in many empirical applications (see for example Al-Rabbaie and Hunt, 2004; Aslan, 2008).

3.3.2 The Hadri test

The test proposed by Hadri (2000) is a residual-based Lagrange Multiplier test (LM) in the spirit of the KPSS test suggested by Kwiatkowski et al. (1992). The Hadri test has a reverse null hypothesis, i.e. that the time series for each cross-section unit is stationary around a deterministic level or trend, against the alternative hypothesis of a unit root. It is based on the following regression:

$$y_{it} = \alpha_i + \gamma_i t + \sum_{t=1}^T u_{it} + \varepsilon_{it} \quad (18)$$

where α_i and $\gamma_i t$ are the deterministic terms defined in (16) above, and the error term has two components: ε_{it} which is white noise, and $\sum_{t=1}^T u_{it}$ which is a random walk. Under the null hypothesis $H_0 : \sigma_u^2 = 0$, whereas under the alternative hypothesis $H_1 : \sigma_u^2 > 0$. That is the null hypothesis is that the variance of the random walk component (σ_u^2) is zero. The autocorrelation in ε_{it} is accounted for by considering the long-run variances of the ε_{it} which is estimated by

$$\hat{s}_{\varepsilon_i}^2 = \frac{1}{T-1} \sum_{t=2}^T \hat{\varepsilon}_{it}^2 + 2 \sum_{j=1}^{\bar{k}} w_{\bar{k}j} \left(\frac{1}{T-1} \sum_{t=j+2}^T \hat{\varepsilon}_{it} \hat{\varepsilon}_{i,t-j} \right) \quad (19)$$

where $w_{\bar{k}j}$ are weights used to ensure that $\hat{\sigma}_{\varepsilon_i}^2$ are always positive. These Bartlett weights are given by $w_{\bar{k}j} = 1 - \left(\frac{j}{\bar{k} + 1} \right)$, where \bar{k} is the bandwidth. The \bar{k} has to be efficiently chosen in order to get reasonable results. The statistic H_0 is the average of the individual KPSS statistic for each series. The test statistic is standardized to take the form $\frac{\sigma_u^2}{\sigma_\varepsilon^2}$, which has a standard normal distribution under the null hypothesis.

3.4 The Pooled-Mean Group Estimator

The test for panel ARDL as proposed by Pesaran et al. (1999) is based on the pooled mean group estimator (PMGE). This allows for the short-run coefficients and error variances to differ across group, but constrains the long-run coefficients to be identical in an error correction framework. The pooled mean group estimator of order (p_i, q_i) is given by the following equation

$$\Delta y_{it} = \phi_i y_{it-1} + \beta_i' x_{it} + \sum_{j=1}^{p_i-1} \lambda_{ij} \Delta y_{i,t-j} + \sum_{j=0}^{q_i-1} \delta_{ij}' \Delta x_{i,t-j} + \alpha_i + \varepsilon_{it} \quad (20)$$

where y_{it} is the dependent variable, x_{it} is vector of explanatory variables, α_i are country specific intercepts and λ_{ij} and δ_{ij} are the country specific coefficients of the short-term dynamics, ε_{it} the

white noise error term. The long-run coefficients $\phi_i = \phi$ (are defined to be the same across countries). If $\phi_i < 0$, there is a long-run relationship between y_{it} and x_{it} defined by $y_{it} = -(\beta'_i / \phi_i) x_{it} + \eta_{it}$. The maximum likelihood procedure based on the concentrated likelihood function is used to estimate the PMGE. A Gauss-Newton algorithm is used to maximize the likelihood function. The pooled mean group regression equation (18) can be estimated with individual specific parameters ϕ_i which are then averaged over N to obtain a PMGE.

3.5 Linear Programming Formulation

The general formulation of a linear programming problem can be stated as follows:

Maximize or minimize the objective function

$$z = CX^T \quad (21)$$

subject to $AX \geq \text{or} \leq b, b \geq 0, X \geq 0$.

where $X = (x_1, x_2, \dots, x_n, x_{n+1}, \dots, x_{n+m})$,

$$C = (c_1, c_2, \dots, c_n, \mathbf{0}, \mathbf{0}, \dots, \mathbf{0}) \text{ and } b = (b_1, b_2, \dots, b_m),$$

$$A = \begin{pmatrix} a_{11} & a_{12} & \dots & a_{1n} & \mathbf{1} & \mathbf{0} & \dots & \mathbf{0} \\ a_{21} & a_{22} & \dots & a_{2n} & \mathbf{0} & \mathbf{1} & \dots & \mathbf{0} \\ \vdots & \vdots & \ddots & \vdots & \vdots & \vdots & \ddots & \vdots \\ a_{m1} & a_{m2} & \dots & a_{mn} & \mathbf{0} & \mathbf{0} & \dots & \mathbf{1} \end{pmatrix}.$$

The model formulated in equation (13) can therefore be expressed as a linear programming problem as follows:

Minimize the objective function

$$m_t = \phi_1 d_t + \phi_2 b_t + \phi_3 \pi_t \quad (22)$$

subject to $\pi_t \leq \mathbf{0.1}$

$$b_t \leq \mathbf{0.04} \quad (23)$$

$$d_t \leq 0.1$$

$$m_t \geq 0.25.$$

The optimal solution is obtained by the use of the Simplex method as shown in Hillier and Lieberman (2001).

4. RESULTS AND DISCUSSION

A preliminary step is taken to calculate the correlation matrix for the four variables. The result shows that government external reserve is most correlated with the fiscal deficit to GDP ratio.

$$\begin{matrix} & \pi_t & b_t & d_t & m_t \\ \pi_t & \left(\begin{matrix} 1.0000 \\ 0.1161 & 1.0000 \\ 0.4206 & -0.2259 & 1.0000 \\ -0.2657 & 0.4590 & -0.2846 & 1.0000 \end{matrix} \right) \\ b_t & \\ d_t & \\ m_t & \end{matrix}$$

4.1 Results of Panel Unit Root Tests

The variables are tested for panel unit root and panel stationarity using Im et al. (2003) and Hadri (2000) to determine their degree of integration and level of stationarity. The results of the panel unit root tests are given in Table 3. The result shows that the variables are nonstationary in levels but stationary in first difference, an indication that the variables are integrated of order 1.

Table 3: Panel Unit Root Tests

Variable	IPS-W-test		Hadri-test	
	No trend	Trend	No trend	Trend
m_t	-1.31(0.09)	-0.04(0.48)	2.65(0.00)	4.13(0.00)
Δm_t	-2.76(0.003)	-1.02(0.155)	0.95(0.17)	3.34(0.00)

π_t	-1.19(0.12)	-0.12(0.45)	0.97(0.17)	3.79(0.00)
$\Delta\pi_t$	-1.45(0.07)	-0.28(0.39)	0.16(0.44)	0.12(0.45)
d_t	-1.99(0.02)	-0.55(0.29)	2.53(0.01)	6.68(0.00)
Δd_t	-3.28(0.0005)	-2.57(0.005)	1.25(0.11)	7.82(0.00)
b_t	0.203(0.58)	0.57(0.72)	1.96(0.03)	5.19(0.00)
Δb_t	-0.83(0.204)	0.02(0.51)	2.51(0.01)	13.08(0.00)

The IPS - w - test at 5% are: for N=5, T=50. The numbers in parentheses are the p-values.

4.2 Panel Cointegration Results

The results of the long-run cointegration parameters using the PMGE are given in Table 4a. The long-run panel cointegration test results show that these variables, inflation rate, fiscal deficit/surplus to GDP ratio account for most of the gross external reserves in agreement with the results of Catao & Terrones (2003).

Table 4a: Estimation Results of PMGE

Variables(Dep. var m_t)	PMGE
π_t	0.44(0.10)
d_t	-0.07(0.03)
b_t	0.03(0.12)

The numbers in parentheses are the standard errors

Central Bank financing of the fiscal deficit however, did show little influence on the gross external reserves with a low coefficient value of -0.07 with standard error of 0.03. The long-run panel cointegration relation using the PMGE with individual constants is

$$ec_t^{m_t} = \mathbf{0.44} \pi_t + \mathbf{0.03} b_t - \mathbf{0.07} d_t \quad (24)$$

Equation (24) shows that a 1% increase in government external reserves (m_t) induces 0.44% inflation rate (π_t). Similarly, a 1% increase in m_t will induce a 0.03% increase in the fiscal deficit to GDP ratio, and -0.07% decrease in the Central Bank financing of fiscal deficit. This result compares favorably with Solomon and de Wet (2004) whose simulation study showed that inflation is very responsive to shocks in budget deficit as well as GDP.

In Table 4b, the estimation results for the PMGE for the individual countries are reported. The error correction terms have values that are reasonably substantial for most of the countries except Guinea. The estimated value of the error correction term for Guinea is 0.002 which is low compared to the other countries. This may not necessarily be due to inefficiency of the method, but the fact that in the long-run government external reserve is not affected by inflation, fiscal deficit to GDP ratio and Central bank financing of fiscal deficit in Guinea (see Banerjee et al. 1993 for details).

Table 4b: Estimation Results of PMGE for Individual Countries

Variables	Gambia	Ghana	Guinea	Nigeria	Sierra Leone
ec_t	-1.35(0.4)	-0.69(0.12)	0.002(0.07)	-0.05(0.26)	-0.39(0.20)
$cons$	5.14(2.20)	-1.19(1.25)	-0.24(0.53)	2.56(2.10)	0.41(0.82)
π_t	-0.28(0.12)	-0.20(0.01)	-0.08(0.45)	-0.02(0.32)	-0.12(0.06)
d_t	0.09(0.03)	-0.01(0.004)	0.02(0.01)	-0.12(0.13)	0.01(0.03)
b_t	0.42(0.15)	-0.30(0.06)	0.02(0.11)	-0.36(1.004)	0.14(0.12)

The numbers in parentheses are the standard errors.

4.3 The Optimum Convergence Criteria

The objective value of the optimum convergence criteria is 0.0452. The objective value for the variables, objective coefficients and their objective value contribution are given in Table 5. Inflation is the highest contributor to the optimum value followed by Central Bank financing of fiscal deposit, while fiscal deficit to GDP ratio is the least.

Table 5: The Optimum Convergence Criteria

Variables	Value	Objective Coefficient	Objective Value Contribution	Slack-/Surplus+
m_t	0.25	0.00	0.0000	0.0000
π_t	0.10	0.44	0.0440	0.0000
d_t	0.00	0.03	0.0012	0.0000
b_t	0.04	-0.07	0.0000	0.1000

5. CONCLUSION

The conclusion we draw from the paper is that Government external reserves, inflation rate, fiscal deficit/GDP ratio, Central Bank financing of fiscal deficit are correlated with the highest correlation between government external reserves and fiscal deficit and GDP ratio. The result shows that the objective value of 0.0462 is obtained with inflation contributing more to the variation in the government external reserves at variance with the correlation matrix. This may be due to the non-stationarity of the variables in level. Due to the result of linear programming approach, we advise that Central Banks in the countries studied to be cautious in implementing inflation targeting as a way of tackling their economic problems. In conclusion there is a warning by Solomon and De Wet (2004) that governments should note the sensitivity of price levels to fiscal policy.

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